

Is It Possible to Know Cosmological Fine-tuning?

Daniel Andrés Díaz-Pachón¹, Ola Hössjer², and Calvin Mathew³ ¹University of Miami, 1120 NW 14 St., Suite 1057, Miami, FL 33136, USA ²Stockholm University, SE-106 910, Stockholm, Sweden ³American Heritage School, 12200 W Broward Blvd., Broward, FL 33325, USA

Received 2024 January 26; revised 2024 February 21; accepted 2024 February 21; published 2024 April 5

Abstract

Fine-tuning studies whether some physical parameters, or relevant ratios between them, are located within socalled life-permitting intervals of small probability outside of which carbon-based life would not be possible. Recent developments have found estimates of these probabilities that circumvent previous concerns of measurability and selection bias. However, the question remains whether fine-tuning can indeed be known. Using a mathematization of the concepts of learning and knowledge acquisition, we argue that most examples that have been touted as fine-tuned cannot be formally assessed as such. Nevertheless, fine-tuning can be known when the physical parameter is seen as a random variable and it is supported in the nonnegative real line, provided the size of the life-permitting interval is small in relation to the observed value of the parameter.

Unified Astronomy Thesaurus concepts: Anthropic principle (48); Bayesian statistics (1900); Analytical mathematics (38); Cosmological constant (334); Cosmological parameters (339)

1. Introduction

Cosmological fine-tuning (FT) says that some physical parameters, some ratios between them, and some boundary conditions must pertain to intervals of small probability to permit the existence of carbon-based life (Lewis & Barnes 2016). FT started to spark the interest of the scientific community when Carter (1974) brought it to light, calling it the anthropic principle, a term that, according to Lewis & Barnes (2016), was popularized and modified by Barrow & Tipler (1988). Since then, considerable effort has been made to determine the length of these so-called life-permitting intervals (LPIs), constituting the area on which the scientific literature has primarily focused (see, e.g., Davies 1982; Barnes 2012; Adams 2019). For this reason, some have linked FT with the small lengths of the intervals, with no regard for their probability (see, e.g., Helbig 2023). This evokes the idea of "naturalness" in physics, where a good theory is assumed to present numbers that are close to unity, and departures from it point to ad hoc explanations trying to fit observations to theory, as Mercury epicycles in the Ptolemaic model (Hossenfelder 2020). FT goes beyond naturalness, asserting that, according to our current theories, such unnaturalness is needed if carbon-based life is going to exist (Barnes 2021). The existence of life sets a specification; i.e., a subset of possible outcomes of parameters for which life is permitted. These outcomes maximize a function f that quantifies how specified outcomes are, and it is stochastically independent of the original random variable. In cosmological FT, this random variable is a physical parameter, whose value is assumed to be independent of the requisites for carbon-based life existence, and the LPI maximizes an indicator function f that determines if life is possible or not. Therefore, the tuning of a physical constant for life is fine if and only if the probability of its LPI is small (Díaz-Pachón & Hössjer 2022).

Original content from this work may be used under the terms of the Creative Commons Attribution 4.0 licence. Any further distribution of this work must maintain attribution to the author(s) and the title of the work, journal citation and DOI.

FT divides the scientific and philosophical communities between those who dismiss it as mere speculation (e.g., McGrew et al. 2001; Colyvan et al. 2005; Hossenfelder 2020) and those who consider it a serious scientific endeavor (e.g., Davies 2008; Tegmark 2015; Lewis & Barnes 2016). One of the main criticisms, even among those who see it as a legitimate question, is the lack of a valid probability distribution to impose over the parameters (Adams 2019). In this direction, Hossenfelder asks: "how do you make a statement about probability without a probability distribution?" (Hossenfelder 2020, p. 205). The problem is indeed daunting: it requires estimating the probability of the LPI using a biased sample of size 1 taken from an unknown distribution supported in an unknown space. More explicitly, and to establish some notation, let \mathcal{X} be the possible values of a random physical parameter X. FT is about inferring the probability of its LPI ℓ_X when the only observation we possess is x_0 , the observed value of X in our Universe; moreover, such an observation suffers from selection bias because that Universe harbors life; and the only information about the probability distribution $F = F_X$ of X is that F is supported on \mathcal{X} and $x_0 \in \mathcal{X}$. Moreover, assuming Carter's definition, it is just natural to assume that F corresponds to a prior distribution of Xthat maximizes the entropy (Carter 1974).

Past efforts to find the probabilities of LPIs use mainly uniform distributions over the set of possible values that constants might take (see, e.g., Tegmark et al. 2006; Barnes 2019–2020; Sandora 2019a). Nonetheless, this approach attests to the difficulty of the task, since the choice of the uniform distribution is based on Bernoulli's Principle of Insufficient Reason (Bernoulli 1713), which requires knowledge that the sample space is finite—a strong assumption that cannot be warranted for cosmological FT. This legitimate criticism of FT measurements came from philosophy, where it was called the normalization problem (McGrew et al. 2001, 2018; Colyvan et al. 2005; McGrew & McGrew 2005).

Thus, the tuning problem can be divided into two steps. First, determining the size of the LPI for a given constant of nature—a task of physics; and second, determining the probability of the LPI—a mathematical task. For this reason, we formally define FT next. **Definition 1** (Fine-tuning). FT happens if and only if $F_0(\ell_X)$ is small, where $F_0 = F_X$ is the assumed true distribution of X, which is supported on \mathcal{X} . In more detail, there is FT if and only if there exists a real-valued $\delta > 0$ such that

$$F_0(\ell_X) \leqslant \delta \ll 1. \tag{1}$$

If the tuning is not fine, we will call it coarse.

Recently, Díaz-Pachón, Hössjer, and Marks developed Algorithm 1 to find an upper bound on the probability of the known LPI ℓ_X of a parameter *X* (Díaz-Pachón et al. 2021, 2023). Algorithm 1 can be viewed as a statistical test, where the null hypothesis H_0 that the tuning of *X* is coarse is tested against the alternative hypothesis H_1 that *X* is fine-tuned. The first two steps of Algorithm 1 ask for the input: a parameter space \mathcal{X} (e.g., \mathbb{N} , \mathbb{R}^+ , \mathbb{R} , \mathbb{R}^n , \mathbb{C} , or some subset thereof) and the constraints to impose on the prior distributions F_X supported on \mathcal{X} (e.g., if some event $B \subset \mathcal{X}$ has a known probability, if some of the moments of *X* are finite, etc.). In the third step, a collection

$$\mathcal{F} := \{ F(\cdot; \theta); \, \theta \in \Theta \}$$

of all the possible maximum entropy (maxent) priors F is determined that satisfy these constraints. These priors are indexed by a hyperparameter θ that varies over a hyperparameter space Θ . In particular, we will assume that the true distribution of X satisfies

$$F_0 = F(\cdot; \theta_0) \in \mathcal{F} \tag{2}$$

for some $\theta_0 \in \Theta$. In the fourth step of Algorithm 1, the maximum probability TP_{max} of the LPI ℓ_X is found over the collection of all the posterior probabilities $\{F(\ell_X; \theta); \theta \in \Theta\}$ induced by the priors in \mathcal{F} . That is, for each $\theta \in \Theta$, we first define the corresponding marginal probability of ℓ_X as

$$F(\ell_X; \theta) = \int_{\mathcal{X}} P(\ell_X | x) F(dx; \theta).$$

by integrating over all possible outcomes *x* of $X \sim F(\cdot; \theta)$. We refer to $F(\ell_X; \theta)$ as a tuning probability. By maximizing this tuning probability over Θ , the maximum tuning probability

$$TP_{\max} := \max_{\theta \in \Theta} F(\ell_X; \, \theta)$$

is obtained. In the fifth step, the outcome of Algorithm 1 is produced: if the test statistic TP_{max} is small ($\leq \delta$ according to Definition 1), we reject the null hypothesis H_0 and conclude there is FT. If TP_{max} is not small, the algorithm is inconclusive. In this case, we cannot reject H_0 . This does not mean that X is not fine-tuned, though. Indeed, H_1 may still be true when H_0 is not rejected, since $F_0(\ell_X)$ may be larger or smaller than δ when $TP_{\text{max}} > \delta$.

Algorithm 1. Algorithm for Testing FT

Input: choose the set of possible values \mathcal{X} of X.

- **Input:** define constraints on the distribution $F = F_X$ of X.
- Find the family $\mathcal{F} = \{F(\cdot;\theta); \theta \in \Theta\}$ of maxent distributions F with support \mathcal{X} , subject to the constraints in Step 2, and assume (A) that the true distribution of X satisfies $F_0 = F(\cdot;\theta_0) \in \mathcal{F}$ for some $\theta_0 \in \Theta$.
- Find the maximum tuning probability $TP_{\max} := \max \{F(\ell_X); F \in \mathcal{F}\}$ over \mathcal{F} for the LPI ℓ_X .
- **Output:** if TP_{max} is small ($\leqslant \delta$), there is FT (reject the assumption that Equation (1) is false; H_0); otherwise, the test is inconclusive (do not reject H_0).

 Table 1

 Maximal Tuning Probabilities TPmax

x	${\mathcal F}$	Θ	Constraint	TP _{max}
	Scale	\mathbb{R}^+	$\epsilon \ll 1$	$2\epsilon C_1$
\mathbb{R}^+	Form and	$\mathbb{R}^+\times\mathbb{R}^+$	None	1
	scale			
	Form and	$\mathbb{R}^+\times\mathbb{R}^+$	$S/N \leq T, T \gg 1, \epsilon \ll 1/\sqrt{T}$	$2\epsilon C_2 \sqrt{T}$
	scale			
	Scale	\mathbb{R}^+	$0 \notin \ell_X, \epsilon \ll 1$	$2\epsilon C_1$
	Scale	\mathbb{R}^+	$0 \in \ell_X$	1
\mathbb{R}	Location	\mathbb{R}	$\epsilon \ll \min(1/C_3, 1)$	$2\epsilon C_3$
	Location and scale	$\mathbb{R}\times\mathbb{R}^+$	$S/N \leq T, \epsilon \ll \min(1/\sqrt{T}, 1)$	$2\epsilon (C_3\sqrt{T}+C_1)$
	Location and scale	$\mathbb{R}\times\mathbb{R}^+$	None	1

Algorithm 1 has some important properties. First, the upper bound TP_{max} is sharp because if the family \mathcal{F} is made of a single maxent distribution F_0 , then $TP_{\text{max}} = F_0(\ell_X)$. For instance, assume that Step 1 selects $\mathcal{X} = \mathbb{R}^+$, and in Step 2 it is assumed that the first moment is finite and known (e.g., $E_0X = \theta_0$). Then the family of Step 3 is made of a single distribution with density $F'(x; \theta_0) = f(x; \theta_0) = e^{-x/\theta_0}/\theta_0$, the exponential with mean θ_0 , since this distribution is maxent among all the distributions with mean θ_0 that are supported in \mathbb{R}^+ . Consequently, $TP_{\text{max}} = F_0(\ell_X) = \int_{\ell_X} e^{-x/\theta_0} dx/\theta_0$.

Second, the selection of the class \mathcal{F} as a family of maxent distributions simultaneously solves both the normalization and selection bias problems (for a detailed discussion, see Díaz-Pachón et al. 2023), maybe the two biggest concerns raised by physicists (Tegmark et al. 2006; Adams 2019; Hossenfelder 2020, 2021) and philosophers (McGrew et al. 2001; Bostrom 2002; Colyvan et al. 2005) alike. On one side, considering more general classes of maxent distributions than uniform ones solves the normalization problem. At the same time, by considering the whole class \mathcal{F} , and not only the distribution induced by an estimate of θ_0 , from the observed value of X in our Universe, the selection bias problem is solved. Consider, for instance, the case when $\mathcal{X} = \mathbb{R}^+$ and \mathcal{F} is the class of exponential distributions with expected value $\theta \in \mathbb{R}^+$. It can then be shown that the maximum likelihood estimator of θ_0 is $\hat{\theta}_0 = X$. But we still have no guarantee that $F_0(\ell_X) = F(\ell_X; \theta_0) \leq$ $F(\ell_X; \hat{\theta}_0)$. On the other hand, by Equation (2) we know that $F_0(\ell_X) \leq TP_{\text{max}}$. Thus, considering all possible values of θ removes the bias induced by the weak anthropic principle, since the class \mathcal{F} includes $F(\cdot, \theta_0)$, but is not reduced to it.

Third, Algorithm 1 reveals that the output heavily depends on the input, namely the sample space \mathcal{X} for X and the constraints to impose on the family \mathcal{F} of prior distributions Fof X. Moreover, Díaz-Pachón, Hössjer, and Marks proved a theorem (Appendix 3 of Díaz-Pachón et al. 2023, summarized in Table 1 below and explained in Remark 1) where they show that subtle changes in the input might produce extremely different outcomes. This can be seen, for instance, from rows 4 to 5 of Table 1, when considering families \mathcal{F} of distributions over $\mathcal{X} = \mathbb{R}$ that depend on a scale parameter only: if $0 \in \ell_X$, it will provoke $TP_{\text{max}} = 1$, regardless of the size of the interval ℓ_X (therefore the level of tuning cannot be assessed), whereas $0 \notin \ell_X$ with a midpoint x_0 of ℓ_X will produce a small TP_{max} , provided the relative half-size $\epsilon = |\ell_X|/(2x_0)$ of the interval is small. In the same direction, when $\mathcal{X} = \mathbb{R}^+$ (= \mathbb{R}) and the form and scale (location and scale) family \mathcal{F} is considered, TP_{max} jumps from very small to 1, depending on whether the maximal signal-to-noise ratio (S/N) T of the prior distributions $F \in \mathcal{F}$ in rows 2–3 (7–8) is bounded or not.

Remark 1. The results of Table 1 are presented for LPIs of the form

$$\ell_X = x_0[1 - \epsilon, 1 + \epsilon], \tag{3}$$

where ϵ is a dimensionless quantity representing the relative half-size of the interval and x_0 is the observed value of X, taken to be the middle point of the interval (if x_0 does not coincide with the middle point of ℓ_X , nothing is lost by assuming the two values coincide and the notation is greatly simplified). When $\epsilon > 0$ is small, FT probabilities are computed as a function of ϵ for diverse parametric families \mathcal{F} of prior distributions F of a randomly generated universe X. These TP_{max} are presented given certain constraints on ℓ_X and/or $\theta \in \Theta$, the latter of which includes one or two variable hyperparameters of the prior densities $f(x; \theta) = F'(x; \theta)$ in \mathcal{F} . The different choices of \mathcal{F} correspond to a prior density $f(x/\theta)/\theta$ for the scale family with scale parameter θ , a prior density $f(x/\theta_2; \theta_1)/\theta_2$ for the form and scale parameter with form parameter θ_1 and scale parameter θ_2 , and a prior density $f((x - \theta_1)/\theta_2)/\theta_2$ for the location and scale family with location parameter θ_1 and scale parameter θ_2 . When \mathcal{F} involves two hyperparameters, the constraints on these hyperparameters are formulated in terms of the maximal value T of the signal-to-noise-ratio S/N = $E^{2}(X)/Var(X)$, that is, the maximal value of the ratio of the squared first moment and variance of the distribution F. There is also a constant associated to each family: \mathcal{F} ; $C_1 =$ $\max_{x>0} xf(x)$ for the scale family; $C_2 = 1/\sqrt{2\pi}$ for the form and scale family; and $C_3 = \max_{x \in \mathbb{R}} f(x)$ for the location and location scale families.

2. Mathematical Framework for Learning and Knowledge Acquisition

The No Free Lunch theorems (Wolpert & Mac-Ready 1995, 1997) assert that, on average, no search does better than a blind one, and therefore a guided search infuses information to the search problem. Active information was thus defined to measure the amount of information introduced by a programmer in order to reach a target, compared to a baseline distribution that is usually, but not necessarily, in maxent (Dembski & Marks 2009; Díaz-Pachón & Marks 2020a). Formally, active information is defined as

$$I^+(A) := \log \frac{\boldsymbol{P}(A)}{\boldsymbol{P}_0(A)},\tag{4}$$

where the target $A \subset \Omega$ is a subset of a search space Ω , whereas P and P_0 are probability distributions on Ω that represent searches of the programmer and blind search, respectively.

Since its inception, active information has been used in the measurement of bias for machine-learning algorithms (Montañez 2017a, 2017b; Montañez et al. 2019, 2021), hypothesis testing

(Díaz-Pachón et al. 2020; Hom et al. 2023), statistical genetics (Díaz-Pachón & Marks 2020b; Thorvaldsen & Hössjer 2023), bump hunting (Díaz-Pachón et al. 2019; Liu et al. 2023), and the estimation and correction of prevalence estimators of Covid-19 (Hössjer et al. 2023; Zhou et al. 2023), among others. In fact, active information can be used as a measure of FT if the search space Ω equals the sample space \mathcal{X} of the physical parameter X, Equation (4) is large for $A = \ell_X$, with $P_0(A) = F_0(A)$ as in Equation (1) and $P(A) = \delta_{x_0}(A)$ a one-point distribution at x_0 , the observed value of X (or the midpoint of ℓ_X ; Thorvaldsen & Hössjer 2020; Díaz-Pachón & Hössjer 2022).

Based also on active information, a mathematical formalization of the epistemological notions of learning and knowledge acquisition was recently developed by Hössjer et al. (2022), where knowledge is usually defined as "justified true belief" (Gettier 1963; Ichikawa & Steup 2018; Schwitzgebel 2021). This means that a subject or agent *S* knows a proposition *p* if

i. S believes p,

- ii. p is true,
- iii. S's belief about p is justified.

If only properties *i* and *ii* are satisfied, we say that *S* learns *p*. In other words, there is learning if and only if there is a true belief. Thus, there is a hierarchization from belief to knowledge through learning:

knowledge
$$\subset$$
 learning \subset belief.

Moreover, the inclusions are proper, since it is possible to have a belief in a false proposition so that such a belief does not constitute learning; and it is also possible to learn p without getting to know p, if the true belief cannot be justified.

As for the mathematical formalization, belief is defined as a probability, as is customary in Bayesian theory (see, e.g., MacKay 2003; Berger 2010), whereas propositions are parameters that have a true value. More explicitly, suppose we want to learn whether a proposition p is either true or false. A set of possible worlds Ω is defined, where one world ω_0 represents the value of the parameter ω , referred to as the true world, whereas all other $\omega \in \Omega \setminus \{\omega_0\}$ are counterfactuals. The set A of interest is made of the worlds in which the proposition p is true. An ignorant person assigns beliefs to every subset of Ω according to some initial distribution P_0 , whereas an agent S with some data D and discernment \mathcal{G} (corresponding to a σ -algebra generated by the nontrivial events that data cannot discern into smaller events) updates his beliefs to P, according to Bayes's rule,

$$\boldsymbol{P}(A) := \frac{L(D|A)\boldsymbol{P}_0(A)}{L(D)},\tag{5}$$

with L(D|A) the likelihood of observing D given A. Then the learning of p is defined as follows.

Definition 2 (Learning). (I) Agent *S* has learned about proposition *p*, compared to an ignorant person, either when *p* is true and the posterior belief *P* about *p* is higher than the prior belief *P*₀ about *p*, or when *p* is false and the posterior belief about *p* is smaller than the prior belief about *p*. (II) The agent *S* has fully learned *p* (regardless of the beliefs of the ignorant person) if the posterior belief *P* about *p* is 1 (0) when *p* is true (false).

Remark 2. Mathematically, the two parts of Definition 2 can be phrased as follows. (I) We say that there is learning about p, compared to an ignorant person, if

$$\begin{cases} 0 < I^+(A), & \text{and } p \text{ is true in the true world } \omega_0, \\ 0 > I^+(A), & \text{and } p \text{ is false in the true world } \omega_0. \end{cases}$$
(6)

(II) There is full learning about p (regardless of the beliefs of the ignorant person) if P(A) = 1 (P(A) = 0) when p is true (false) in the true world ω_0 .

In particular, a subject in a maximum state of ignorance is represented by a maxent distribution P_0 over Ω .

Nonetheless, learning does not necessarily entail a particular belief about the true world, so it does not satisfy the conditions of a justified true belief, which requires having a belief for the right reasons. Knowledge acquisition is defined to cover this gap.

Definition 3 (Knowledge). (I) Agent S has acquired knowledge about p, compared to an ignorant person, if the following three conditions are satisfied:

- 1. S has learned about p, compared to the ignorant person.
- 2. The true world ω_0 is among the pool of possibilities for *S*, according to his posterior beliefs.
- 3. The belief in ω_0 under **P** is stronger than that under **P**₀.

(II) Agent *S* has acquired full knowledge about *p* (regardless of the beliefs of the ignorant person) if $P = \delta_{\omega_0}$.

Remark 3. Mathematically, the three requirements of knowledge acquisition for Part (I) of Definition 3 amount respectively to:

- 1. The criteria of Part (I) of Definition 2 are satisfied.
- 2. The true world ω_0 is in supp(**P**), the support of **P**.
- 3. For all $\epsilon > 0$, the closed ball $B_{\epsilon}[\omega_0] := \{\omega \in \Omega: d(\omega, \omega_0) \leq \epsilon\}$ is such that $I^+(B_{\epsilon}[\omega_0]) \geq 0$, with strict inequality for some $\epsilon > 0$, where *d* is a metric over Ω .

Condition 1 ensures that knowledge is a more stringent concept than learning. Condition 2 is mathematically equivalent to saying that *S* has a positive belief for every open ball centered at ω_0 (i.e., if for all $\epsilon > 0$, $P(B_{\epsilon}(\omega_0)) > 0$, where $B_{\epsilon}(\omega_0)$: ={ $\omega \in \Omega$: $d(\omega, \omega_0) < \epsilon$ } is the open ball of radius ϵ centered at ω_0), which in turn explains Condition 3.

3. Learning and Knowledge in Cosmology

It is pointless to think of different universes with different laws of nature because they are beyond our comprehension. Therefore, following the reasonable little question of Barnes (2019–2020), we reduce the set of possible universes to those with the same laws but possibly different constants. In such a scenario, there is a one-to-one relation between \mathcal{X} and the multiverse, because each $x \in \mathcal{X}$ represents a universe with value *x* for the parameter *X*. In other words, there is a bijective function from \mathcal{X} to the set of possible universes (Adams 2019; Sandora 2019a, 2019b, 2019c, 2019d; Sandora et al. 2022, 2023a, 2023b, 2023c), with x_0 being the actual value of the constant of nature in our world, and any other $x \in \mathcal{X}$ being a counterfactual universe. The framework from Section 2 will be applied to analyze conditions under which FT can be known. As a preparation, we first consider the simpler proposition that our Universe harbors life (which we know to be true).

Theorem 1. The proposition

p_1 : Our Universe harbors life

can be fully learned and known. A sufficient condition for learning and knowledge acquisition of p_1 , compared to an ignorant person, is having $TP_{\text{max}} < 1$.

Proof 1. Identify the set of possible worlds Ω for the learning and knowledge acquisition problem with the set of possible universes \mathcal{X} , with ω_0 corresponding to the observed value x_0 of X for our Universe. According to Section 2, the set $A = A_1$ of worlds for which p_1 is true coincides with ℓ_X , the LPI of X.

Let P be given by our best current theories and data. Since our theories and data admit knowledge of $\omega_0 = x_0$ (i.e., we know that our Universe admits life as well as the observed value x_0 of X), it follows that $P = \delta_{x_0}$ has a one-point distribution at x_0 . Since $x_0 \in \ell_X$, this implies

$$\boldsymbol{P}(A_1) = \boldsymbol{P}(\ell_X) = 1. \tag{7}$$

Then, according to Part (II) of Definition 2 and Remark 2, we fully learn p_1 because P(A) = 1 and p_1 is true. Since $P = \delta_{x_0}$, it follows from Part (II) of Definition 3 that we also acquire full knowledge about p_1 . This completes the proof of the first part of Theorem 1.

For a proof of the second part of Theorem 1, we also need to consider the beliefs P_0 of the ignorant person about the value of X. Recall that P_0 is in maxent over $\Omega = \mathcal{X}$, given the restrictions imposed by \mathcal{F} , the class of distributions considered in Table 1. This implies $P_0 \in \mathcal{F}$ and consequently

$$\boldsymbol{P}_0(A_1) = \boldsymbol{P}_0(\ell_X) \leqslant \max_{F \in \mathcal{F}} F(\ell_X) = TP_{\max}$$

Assume $TP_{\text{max}} < 1$. From this assumption and Equation (2), it follows that $P_0(A_1) < P(A_1)$, which, according to Part (I) of Definition 2 and Remark 2, implies learning about p_1 compared to an ignorant person, since p_1 is true. The assumption $TP_{\text{max}} < 1$ also implies knowledge acquisition about p_1 compared to an ignorant person, since Conditions 2 and 3 of Definition 3 are satisfied as well.

Remark 4. Notice that p_1 does not include the concept of FT.

After having dealt with learning and knowledge acquisition of our Universe harboring life, let us now study whether it is possible to learn and know that our Universe is fine-tuned. This corresponds to the proposition

 p_2 : Our Universe is fine-tuned.

We prove the following result.

Theorem 2. Let F_0 be the actual distribution of X and $0 < \delta \ll 1$ the upper-bound probability for FT, as stated in Definition 1. Assume the agent S believes that the assumption

$$F_0 \in \mathcal{F}$$
 (A)

in the third step of Algorithm 1 is true. The active information for the set of worlds A for which p_2 is true satisfies

$$I^{+}(A) = \min(\log(TP_{\max}^{-1}), \log(\delta^{-1})),$$
(8)

from which

follows.

$$I^+(A) \leq -\log TP_{\max}$$

Proof 2. Let $\Omega = (0, 1]$ be the set of possible values of $F_0(\ell_X)$, whereas $\omega_0 \in \Omega$ is the actual value of $F_0(\ell_X)$. Moreover, notice that $0 \leq F_0(\ell_X) \leq TP_{\text{max}}$ because of Equation (A). It follows from Definition 1 that $A = A_2 = (0, \delta]$ is the set of possible values of $F_0(\ell_X)$ that corresponds to a fine-tuned universe, that is, the set of values of $F_0(\ell_X)$ for which p_2 is true.

Since Ω is a bounded set, the maxent prior distribution $P_0 \sim U(0, 1)$ is uniform. Consequently, $P_0(A_2) = \delta$. Interpreting the agent's belief in Equation (A) as new data *D*, and since P_0 is uniform, it follows from Equation (5) that the posterior distribution $P \sim U(0, TP_{\text{max}})$ is uniform as well. Therefore, $P(A_2) = \min(\delta/TP_{\text{max}}, 1)$. From this, it follows that

$$I^{+}(A_{2}) = \log \frac{\min(\delta/TP_{\max}, 1)}{\delta} = \operatorname{logmin}(TP_{\max}^{-1}, \delta^{-1}).$$

Corollary 2.1. Assume that p_2 is true, so that Equation (1) is satisfied for some $0 < \delta \ll 1$. Then

$$TP_{\max} < 1$$
 (9)

is enough to learn p_2 , that X is fine-tuned for life, whereas knowing p_2 requires more. For instance, Equation (A) is a sufficient additional condition to assure that knowledge about p_2 has been acquired.

Proof 2.1. Since p_2 is true by assumption, learning follows from Part (I) of Definition 2 and the fact that $I^+(A) > 0$, which follows from Equations (8), (9), and $\delta < 1$. Note that although the agent's belief in Equation (A) was used to construct the posterior belief P in Theorem 2, it is not required that Equation (A) is true to derive P and the formula in Equation (8) for $I^+(A)$.

If only Equations (1) and (9), but not Equation (A), hold, then it may occur that $TP_{\max} < \omega_0 \leq \delta$, so that for small enough $\epsilon > 0$, we have $0 = P(B_{\epsilon}[\omega_0]) < 2\epsilon = P_0(B_{\epsilon}[\omega_0])$. Whenever this happens, Condition 3 of Definition 3 is violated. On the other hand, if Equations (1), (9), and (A) hold, it is clear that $\omega_0 \in (0, TP_{\max})$. Since $P_0 \sim U(0, 1)$ and $P \sim U(0, TP_{\max})$, it follows from Definition 3 that the agent has acquired knowledge about p_2 compared to the ignorant person.

Remark 5. Note that even though we assumed the agent's belief in Equation (A) to construct P, this condition is not needed for him to learn about FT. Comparing Theorem 1 with Corollary 2.1, corresponding with intuition, we find that more is required to learn that our Universe is fine-tuned (p_2) than to learn that it harbors life (p_1).

Theorem 3. Equation (A) and

$$TP_{\max} \leqslant \delta \ll 1$$
 (B)

are sufficient for establishing, learning, and knowing that X is fine-tuned for life.

Proof 3. Since Equations (A) and (B) imply Equation (1), they are sufficient to establish FT (that p_2 is true). Also,

Equation (B) implies Equation (9). Therefore, according to Corollary 2.1, it follows that the agent has learned and known FT compared to an ignorant person.

Remark 6. As seen in the proof of Corollary 2.1, it is possible to learn the FT of X for the wrong reasons if Equations (1) and (B) hold, but Equation (A) does not. However, this is unsatisfactory, because it does not say anything about the real probability $\omega_0 = F_0(\ell_X)$ of FT. Indeed, it is possible to falsely learn FT if Equation (1) is false while Equation (9) is true. Therefore, if FT cannot be known, learning is not useful. For this reason, we only pay attention to scenarios where FT can be known.

4. On Knowing FT

4.1. The Crucial Choice of \mathcal{F}

Theorem 3 says that it is enough to focus on Equations (A) and (B), since these two equations imply that we can know that X is fine-tuned for life. To analyze whether Equations (A) and (B) are restrictive, recall that the first two steps of Algorithm 1 determine the size of the set \mathcal{F} of prior distributions of X, which represent possible beliefs of an ignorant person about our Universe. It turns out that these two steps are of crucial importance for the validity of Equations (A) and (B), and hence for establishing that our Universe is fine-tuned. The less we assume in Steps 1 and 2 of Algorithm 1 (the larger \mathcal{F} is), the easier it is to satisfy Equation (A), but the more difficult it is to satisfy Equation (B). This has interesting implications for the list of possible choices of \mathcal{F} in Table 1.

First, since the maximal tuning probability TP_{max} is large for rows 2, 5, and 8 of Table 1, Equation (B) is not satisfied and FT can neither be established nor known in these cases.

On the other hand, rows 3, 4, 6, and 7 of Table 1 represent a second scenario where \mathcal{F} is smaller, making it easy to satisfy Equation (B) but more difficult to satisfy Equation (A). For instance, when the S/N of the distribution F_X of X is bounded (rows 3 and 7), it is still possible that Equation (A) does not hold (e.g., in the $\mathcal{X} = \mathbb{R}$ case, if the true maxent distribution F_0 of X is positive, with a very small standard deviation in relation to its mean).

4.2. Random Distributions

For the second scenario of Section 4.1, when \mathcal{F} is chosen so small that Equation (B) holds but Equation (A) does not, one may assume that the distribution $F = F_X$ of X is random. Under appropriate conditions, this leads to a generalization of Theorem 3, whereby the FT of X can be known, not with certainty but still with a high probability.

In more detail, since $F_X = F(\cdot; \theta)$, such an approach leads to a having a random hyperparameter θ , with the true distribution $F_0 = F(\cdot; \theta_0)$ of X corresponding to an instantiated value θ_0 of θ . It follows from the proofs of Corollary 2.1 and Theorem 3 that if Equation (B) holds, we can know that X is fine-tuned, with a high probability 1 - P(E), if the event

$$E = \{F_0(\ell_X) > TP_{\max}\} \subset \{F_0 \notin \mathcal{F}\} = \{\theta_0 \notin \Theta\}$$
(10)

has low probability. (Note that P(E) being small is less stringent than Equation (A) failing with a small probability.) Showing that P(E) is small thus requires choosing the distribution of $F = F(\cdot; \theta)$, for which a distribution G_{θ} of θ is needed. This requires implementing Algorithm 1 on a new class of hyperhyperparameters λ that parameterize the distribution $G_{\theta} = G(\cdot; \lambda)$ of θ , which in turn leads to new restrictions $\lambda \in \Lambda$ on the hyperhyperparameter. We thus end up in an endless loop (Hofstadter 1999) because it requires showing that the hyperparameters θ themselves are tuned to obtain the FT of the physical parameters *X*. In other words, we replace the FT question of the physical parameters θ .

4.3. Bounded Parameter Space

The strategy of Section 4.2, to regard F_X as random, would be great if the parameter space $\mathcal{X} = [a, b]$ of X were finite and Equation (B) was satisfied for the chosen class \mathcal{F} of distributions of X. If there are no asymmetric restrictions on F, it is natural to regard [a, b] as a torus and assume that F is translation invariant in the sense that $F(\cdot)$ and $F(\cdot - x)$ have the same distribution for all x. It can then be proven, as a consequence of Markov's inequality, that the probability of the set in Equation (10) satisfies

$$P(E) = P[F_0(\ell_X) > TP_{\max}] \leqslant \frac{|\ell_X|}{[(b-a)TP_{\max}]},$$

which is small if

$$|\ell_X| \ll TP_{\max}.$$
 (11)

This implies that, for a finite parameter space \mathcal{X} , we can know that X is fine-tuned, with a large probability 1 - P(E), if \mathcal{F} is chosen small enough so that Equation (B) holds, but also large enough so that Equation (11) is valid.

Since a finite parameter space is a favorite scenario for physicists (remember the discussion after Theorem 1), it is worth exploring whether there are actual FT instances in which the parameter space is actually finite. However, the current arguments in favor of this hypothesis are unconvincing (McGrew et al. 2001; Colyvan et al. 2005). Therefore, we need to deal with \mathcal{X} being unbounded, and for such cases FT cannot be known unless some restrictions, such as Equations (A) and (B), are imposed.

4.4. The Real Line as Parameter Space

The argument of Section 4.3, based on random measures, is unfortunately not straightforward to generalize when the parameter space \mathcal{X} of X equals \mathbb{R} . Rows 4 and 5 of Table 1 correspond to a scale family of distributions that typically are symmetric around the origin. However, the scenario of row 4 does not consider the observed value x_0 of the parameter in our Universe as an estimator of a location parameter for F_X , making it plausible to suggest that Equation (A) is violated. Indeed, it is a basic principle to obtain maxent distributions, subjecting them to the available observations (Wainwright & Jordan 2008; Singh & Vishnoi 2014; Celis et al. 2020; Suh 2023). Thus, defining \mathcal{F} without including into it a random variable whose location parameter is given by the observed value of X in our Universe might be too restrictive.

On the other hand, the location scale family with an upper bound *T* of the S/N (row 7 of Table 1) does include distributions with x_0 as a mean parameter. It is possible then to know that *X* is fine-tuned if the true distribution F_0 satisfies $S/N \le T$ (so that Equation (A) holds) and the relative halflength ϵ of the LPI is small enough so that Equation (B) holds. The assumption $S/N \le T$ cannot be derived from the maxent principle, though, and it rather needs some other justification (such as simplicity).

4.5. Positive Line as Parameter Space

When $\mathcal{X} = \mathbb{R}^+$, it is possible to use row 1 of Table 1. Notice that this is a scale family of distributions over \mathbb{R}^+ that includes all the exponential distributions. This is important because the exponential distribution $F(\cdot; \theta)$ is maxent over all distributions restricted to have mean θ . Other maxent distributions over \mathbb{R}^+ are possible under different constraints; however, they are less parsimonious because they require more restrictions. For instance, the χ^2 distribution of X has maxent distribution over all distributions of the nonnegative reals given the restrictions on E(X) and $E(\log X)$ (see, e.g., Park & Bera 2009), thus pertaining to row 2 of Table 1. Hence, the family of exponential distributions $\{F(\cdot;\theta)\}_{\theta \in \mathbb{R}^+}$ is the less restrictive family of maxent distributions over \mathbb{R}^+ , and we can naturally include into it the exponential distribution with a mean value corresponding to the observed value x_0 of the constant in our Universe. Therefore, in this scenario, it is plausible that Equation (A) is satisfied and, whenever Equation (B) holds as well, we can establish and know FT. To give a more explicit formulation of Equation (B) for the family of exponential distributions of X (see row 1 of Table 1), we notice that the maximum tuning probability equals

$$TP_{\max} = 2C_1\epsilon = 2(\max_{x>0} xe^{-x})\epsilon = 2e^{-1}\epsilon$$
 (12)

when $\epsilon > 0$ is small.

By the reasoning above, we have thus established the following important principle.

Principle 1. The FT of the parameter *X* can be known when:

- 1. The constant can only take nonnegative values ($\mathcal{X} = \mathbb{R}^+$).
- 2. The size of the LPI is small relative to the observed value
- of the constant in our Universe $(2e^{-1}\epsilon \leq \delta \ll 1)$.

5. Examples

We consider several examples of physical parameters X to determine whether FT can be known for them. We will only treat cases in which the parameter cannot take negative values and make use of Principle 1. This section does not intend to be exhaustive. It rather intends to show two things. First, that FT can be known in a few instances. Second, that we can fail to know the level of tuning for a parameter if the relative halflength ϵ of its LPI is not small. More bounds for the LPIs can be found, for instance, in Adams (2019) and references therein. In particular, Table 2 of Adams (2019) shows that many LPIs span over several orders of magnitude. This implies that FT cannot be known in those cases, because ϵ is large and therefore (by Table 1) Equation (B) is not satisfied.

5.1. Critical Density of the Universe

According to Davies (1982), the LPI of the critical density of the Universe $\rho_{\rm crit}$ is

$$\ell_{\rho_{\rm crit}} = \rho_{\rm crit,0} [1 - 10^{-60}, 1 + 10^{-60}],$$

where $\rho_{crit,0}$ is the observed critical density. Since the density cannot be negative, $\mathcal{X} = \mathbb{R}^+$. Taking $\epsilon = 10^{-60}$, we notice that $\epsilon \ll 1$. According to row 1 of Table 1, Díaz-Pachón et al. (2023),

and Equation (12), we have $TP_{\text{max}} = 2e^{-1} \times 10^{-60}$ as a small number. Thus, by Principle 1, we know that there is FT.

5.2. Gravitational Force

According to Davies (1982), when observing the ratio X of the gravitational constant G to the contribution from vacuum energy to the cosmological constant Λ_{vac} , the LPI of the gravitational constant is

$$\ell_X = x_0 [1 - 10^{-100}, 1 + 10^{-100}],$$

where x_0 is the observed value of *X*. Since gravitation is attractive, it cannot be negative, and therefore *X* is positive as well (so that $\mathcal{X} = \mathbb{R}^+$). Since $\epsilon = 10^{-100}$, it follows that $\epsilon \ll 1$. Thus, according to row 1 of Table 1, Díaz-Pachón et al. (2021), and Equation (12), the maximal tuning probability is $TP_{\text{max}} = 2e^{-1} \times 10^{-100}$. By Principle 1, we know that there is FT.

5.3. Higgs Vacuum Expectation Value

The Higgs vacuum expectation value v is given by $v_0 = 2 \times 10^{-17} m_{\rm P}$, where $m_{\rm P}$ is the Planck mass. According to Barnes (2012), its LPI (in units of $m_{\rm P}$) is given by

$$\ell_{\nu} = [0.78 \times 10^{-17}, 3.3 \times 10^{-17}] = [a, b].$$
 (13)

If *v* falls below the lower bound *a*, the proton becomes heavier than the neutron, causing hydrogen to be unstable and enabling electron capture reaction (Damour & Donoghue 2008). If *v* exceeds the upper bound *b* of ℓ_v , there is no nuclear binding. The Higgs vacuum expectation value is positive, as it describes the average value of the Higgs field in the vacuum (i.e., $\mathcal{X} = \mathbb{R}^+$). Writing Equation (13) as in Equation (3), we obtain

$$\ell_{\nu} = \frac{a+b}{2} [1-\epsilon, 1+\epsilon]$$

= 2.04 × 10⁻¹⁷ [1 - 0.617, 1 + 0.617]

. 1

Since $\epsilon = 0.617$ is not small, we cannot use Equation (12) to calculate the maximum tuning probability as $TP_{\text{max}} = 2e^{-1} \times 0.617 = 0.454$. Instead, we make use of the fact that for an exponential distribution $F(\cdot; \theta)$, with mean θ , the tuning probability of ℓ_v equals

$$F(\ell_{v}; \theta) = \exp(-a/\theta) - \exp(-b/\theta).$$

This tuning probability is maximized for $\theta = (b - a)/\log(b/a)$, with a maximal value

$$TP_{\max} = \left(\frac{b}{a}\right)^{-a/(b-a)} - \left(\frac{b}{a}\right)^{-b/(b-a)} = 0.488.$$
(14)

Thus, despite the smallness of ℓ_v , FT cannot be known for v because TP_{max} is large and hence Equation (B) is not satisfied.

5.4. Amplitude of Primordial Fluctuations

According to Rees (2000), the LPI of the amplitude Q of the primordial fluctuations is

$$\ell_{Q} = [10^{-6}, 10^{-5}]$$

= Q₀[1 - 0.818, 1 + 0.818]
= [a, b],

with $Q_0 = 5.5 \times 10^{-6}$ being the midpoint of ℓ_Q . Since the amplitude cannot be negative, $\mathcal{X} = \mathbb{R}^+$. Since $\epsilon = 0.818$ is not

small, we follow the procedure of the previous example and calculate the maximal tuning probability as

$$TP_{\max} = \left(\frac{b}{a}\right)^{-a/(b-a)} - \left(\frac{b}{a}\right)^{-b/(b-a)} = 0.697,$$

in agreement with Díaz-Pachón et al. (2021). Since this number is not small, the FT of Q cannot be known.

5.5. Baryon–Photon Ratio

According to Adams (2019), the baryon–photon ratio η has a value $\eta \sim 6 \times 10^{-10}$, and it can sustain life in between the interval

$$\ell_{\eta} = [6 \times 10^{-13}, 6 \times 10^{-7}] = [a, b].$$
(15)

Again, since ϵ is not small, we follow the procedure of the previous two examples to calculate the maximal tuning probability:

$$TP_{\max} = \left(\frac{b}{a}\right)^{-a/(b-a)} - \left(\frac{b}{a}\right)^{-b/(b-a)} = 0.9999852.$$
(16)

Thus, since TP_{max} is not small, the level of tuning of η cannot be known.

Remark 7. It follows from Examples 5.3 to 5.5 that the maximal tuning probability of the LPI $\ell_X = [a, b]$ of $X \in \mathcal{X} = \mathbb{R}^+$, with a class of exponential priors for F_X , is a function

$$TP_{\max} = h(r) = r^{-1/(r-1)} - r^{-r/(r-1)}$$
(17)

of the ratio $r = b/a = (1 + \epsilon)/(1 - \epsilon)$ between the right and left end points of ℓ_X . It can be seen that $h: (1, \infty) \to (0, 1)$ is a strictly increasing function of r, with $\lim_{r\to 1} h(r) =: h(1) = 0$, $h'(1) = e^{-1}$, and $\lim_{r\to\infty} h(r) = 1$. From this, it follows that Equation (12) is a special case of Equation (17), which holds in the limit of small half relative sizes $\epsilon > 0$ of ℓ_X . Moreover, from Theorem 3, a sufficient condition to know about the FT of X on $\mathcal{X} = \mathbb{R}^+$ is

$$\frac{b}{a} \leqslant h^{-1}(\delta). \tag{18}$$

When $\delta > 0$ is small, Equation (18) is essentially equivalent to the condition on $\epsilon > 0$ in the second part of Principle 1.

6. Conclusion

Using the recent mathematization of knowledge acquisition in Hössjer et al. (2022), we formalize Principle 1, which establishes that the FT of a given physical parameter X can be known if at least the following two conditions are satisfied:

- 1. The constant X can only take nonnegative values.
- 2. The size of its LPI is small relative to the observed value x_0 of the constant in our Universe.

This latter proviso is extremely important since, as we saw in the examples, it is possible to have very small intervals whose size relative to x_0 is large. As a result, Equation (B) is not satisfied and the level of tuning cannot be known. In other words, to know that a given constant is fine-tuned, it is the relative half-size of the interval that must be small, not the interval itself. That is, $|\ell_X| \ll x_0$ is required in order to satisfy Equation (B). Our conclusion is sobering in the sense that many things that, in spite of the ado, are touted as fine-tuned cannot be known to be so.

THE ASTROPHYSICAL JOURNAL SUPPLEMENT SERIES, 271:56 (8pp), 2024 April

On the other hand, when Principle 1 is satisfied, we can be sure that FT can be known. However, we also highlight that not satisfying Equation (B) does not imply that the tuning is coarse; it only implies that FT cannot be known, even if Equation (A) is satisfied, since it is still possible that the real probability $F_0(\ell_X)$ of the tuning is small. Thus, when the maximal tuning probability TP_{max} is not small, the conclusion of coarse tuning is epistemological, not ontological.

ORCID iDs

Daniel Andrés Díaz-Pachón b https://orcid.org/0000-0001-6281-1720

Ola Hössjer () https://orcid.org/0000-0003-2767-8818

References

- Adams, F. C. 2019, PhR, 807, 1
- Barnes, L. A. 2012, PASA, 29, 529
- Barnes, L. A. 2021, in The Routledge Companion to Philosophy of Physics, ed. E. Knox & A. Wilson (Routledge: New York)
- Barnes, L. A. 2019-2020, Ergo, 6, 1220
- Barrow, J. D., & Tipler, F. J. 1988, The Anthropic Cosmological Principle (Oxford: Oxford Univ. Press)
- Berger, J. 2010, Statistical Decision Theory and Bayesian Analysis (2nd ed.; Berlin: Springer)
- Bernoulli, J. 1713, Ars Conjectandi (Basel: Thurneysen Brothers)
- Bostrom, N. 2002, Anthropic Bias: Observation Selection Effects in Science and Philosophy (New York: Routledge)
- Carter, B. 1974, in IAU Symp. 63, Confrontation of Cosmological Theories with Observational Data, ed. M. S. Longhair (Dordrecht: D. Reidel), 291
- Celis, L. E., Keswani, V., & Vishnoi, N. K. 2020, in Proc. Machine Learning Research 119, ICML'20: Proc. of the 37th Int. Conf. on Machine Learning, ed. H. Daumé, III & A. Singh, (PMLR), 1349, https://proceedings.mlr. press/v119/celis20a.html
- Colyvan, M., Garfield, J. L., & Priest, G. 2005, Synthese, 145, 325
- Damour, T., & Donoghue, J. F. 2008, PhRvD, 78, 014014
- Davies, P. 1982, The Accidental Universe (Cambridge: Cambridge Univ. Press) Davies, P. 2008, The Goldilocks Enigma: Why Is the Universe Just Right for
- Life? (New York: Mariner Books)
- Dembski, W. A., & Marks, I. I. 2009, ITSMC, 5, 1051
- Díaz-Pachón, D. A., & Hössjer, O. 2022, Entrp, 24, 1323
- Díaz-Pachón, D. A., Hössjer, O., & Marks, R. J., II 2021, JCAP, 2021, 020
- Díaz-Pachón, D. A., Hössjer, O., & Marks, R. J., II 2023, FoPh, 53, 1
- Díaz-Pachón, D. A., & Marks, R. J., II 2020a, BIO-Complexity, 2020, 1 Díaz-Pachón, D. A., & Marks, R. J., II 2020b, BIO-Complexity, 2020, 3
- Díaz-Pachón, D. A., Sáenz, J. P., Rao, J. S., & Dazard, J.-E. 2019, Appl. Stoch. Models Bus. Ind., 35, 376
- Díaz-Pachón, D. A., Sáenz, J. P., & Rao, J. S. 2020, Stat. Probab. Lett., 161, 108742
- Gettier, E. L. 1963, Analysis, 23, 121
- Helbig, P. 2023, FoPh, 53, 93
- Hofstadter, D. R. 1999, Gödel, Escher, Bach: an Ethernal Golden Braid (New York: Basic Books)
- Hom, C., Yik, W., & Montañez, G. D. 2023, in 2023 IEEE 10th Int. Conf. on Data Science and Advanced Analytics (DSAA) (Piscataway, NJ: IEEE)

- Hossenfelder, S. 2020, Lost in Math: How Beauty Leads Physics Astray (New York: Basic Books)
- Hossenfelder, S. 2021, Synthese, 198, 3727
- Hössjer, O., Díaz-Pachón, D. A., Chen, Z., & Rao, J. S. 2023, ITIT, in press
- Hössjer, O., Díaz-Pachón, D. A., & Rao, J. S. 2022, Entrp, 24, 1469
- Ichikawa, J. J., & Steup, M. 2018, The Analysis of Knowledge, in The Stanford Encyclopedia of Philosophy, ed. E. N. Zalta (Stanford, CA: Metaphysics Research Lab, Stanford Univ.)
- Lewis, G. F., & Barnes, L. A. 2016, A Fortunate Universe: Life In a Finely Tuned Cosmos (Cambridge: Cambridge Univ. Press)
- Liu, T., Díaz-Pachón, D. A., Rao, J. S., & Dazard, J.-E. 2023, ITPAM, 45, 4637
- MacKay, D. J. C. 2003, Information Theory, Inference, and Learning Algorithms (Cambridge: Cambridge Univ. Press)
- McGrew, L., & McGrew, T. 2005, Philosophia Christi, 7, 423
- McGrew, T. 2018, Quaestiones Disputatae, 8, 147
- McGrew, T., McGrew, L., & Vestrup, E. 2001, Mind, 110, 1027
- Montañez, G. D. 2017a, PhD thesis, Carnegie Mellon Univ., Pittsburgh, PA
- Montañez, G. D. 2017b, in 2017 IEEE Int. Conf. on Systems, Man, and Cybernetics (SMC) (Piscataway, NJ: IEEE), 477
- Montañez, G. D., Bashir, D., & Lauw, J. 2021, in Agents and Artificial Intelligence. ICAART 2020, ed. A. P. Rocha, L. Steels, & J. ven den Herik (Berlin: Springer), 332
- Montañez, G. D., Hayase, J., Lauw, J., et al. 2019, in 2nd Australasian Joint Conf. on Artificial Intelligence (AI 2019), ed. J. Liu & J. Bailey (Berlin: Springer), 277
- Park, S. Y., & Bera, A. K. 2009, J. Econom., 150, 219
- Rees, M. J. 2000, Just Six Numbers: The Deep Forces That Shape The Universe (New York: Basic Books)
- Sandora, M. 2019a, Univ, 5, 149
- Sandora, M. 2019b, Univ, 5, 157
- Sandora, M. 2019c, Univ, 5, 171
- Sandora, M. 2019d, Univ, 5, 175
- Sandora, M., Airapetian, V., Barnes, L. A., Lewis, G. F., & Pérez-Rodríguez, I. 2022, Univ, 8, 651
- Sandora, M., Airapetian, V., Barnes, L. A., & Lewis, G. F. 2023a, Univ, 9, 2
- Sandora, M., Airapetian, V., Barnes, L. A., & Lewis, G. F. 2023b, Univ, 9, 4
- Sandora, M., Airapetian, V., Barnes, L. A., Lewis, G. F., & Pérez-Rodríguez, I. 2023c. Univ. 9, 42
- Schwitzgebel, E. 2021, Belief, in The Stanford Encyclopedia of Philosophy, ed. E. N. Zalta (2021 ed.; Stanford, CA: Metaphysics Research Lab, Stanford Univ.)
- Singh, M., & Vishnoi, N. K. 2014, in Proc. of the 46th Annual ACM Symp. on Theory of Computing, ed. D. Shmoys (New York: Association for Computing Machinery), 50
- Suh, C. 2023, Information Theory for Data Science (Norwell, MA/Delft: Now Publishers)
- Tegmark, M. 2015, Our Mathematical Universe (New York: Vintage)
- Tegmark, M., Aguirre, A., Rees, M., & Wilczek, F. 2006, PhRvD, 73, 023505
- Thorvaldsen, S., & Hössjer, O. 2020, JThBi, 501, 110352
- Thorvaldsen, S., & Hössjer, O. 2023, J. R. Stat. Soc. C, 72, 1310
- Wainwright, M. J., & Jordan, M. I. 2008, Found. Trends Mach. Learn., 1, 1
- Wolpert, D. H., & MacReady, W. G. 1995, No Free Lunch Theorems for Search SFI-TR-95-02-010, Santa Fe Institute, https://www.santafe.edu/ research/results/working-papers/no-free-lunch-theorems-for-search
- Wolpert, D. H., & MacReady, W. G. 1997, IEEE Trans. Evol. Comput., 1, 67

Zhou, L., Díaz-Pachón, D. A., Zhao, C., Rao, J. S., & Hössjer, O. 2023, Stat. Med., 42, 4713