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Evolutionary dynamics and fixation probabilities in directed networks

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Abstract. In this paper, we investigate the evolutionary dynamics in directed and/or weighted networks. We study the fixation probability of a mutant in finite populations in stochastic voter-type dynamics for several update rules. The fixation probability is defined as the probability of a newly introduced mutant in a wild-type population taking over the entire population. In contrast to the case of undirected and unweighted networks, the fixation probability of a mutant in directed networks is characterized not only by the degree of the node that the mutant initially invades but also by the global structure of networks. Consequently, the gross connectivity of networks such as small-world property or modularity has a major impact on the fixation probability.

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1. Introduction

Evolutionary dynamics describe the competition among different types of individuals in ecological and social systems. Traits, either genetic or cultural, are transmitted to others through inheritance or imitation. The fitness of an individual determines her/his ability to pass on her/his traits to the next generation. An individual with a larger fitness value is more likely to replace one with a smaller fitness value. Such a dynamical process can be modeled using the well-known voter model and its variants [1]–[5]. According to these models, individuals adopt the trait (hereafter we call it 'type') of others. Selection and random drift are two major driving forces in evolutionary dynamics. Selection results from the different fitness levels of different types. Random drift results from the finite size of populations.

In the voter-type dynamics, one's type is replaced with the type of another individual. Therefore, no new types are introduced into the population unless an explicit mutation (or innovation) is considered. Once a single type dominates the entire population, this unanimity state remains the same forever. In other words, the unanimity states are the absorbing states of these dynamics. Consequently, in the case of finite populations, the stochasticity of voter-type models leads to the fixation or extinction of a newly introduced type after some time. The probability that a single mutant introduced in a population of wild-type individuals eventually takes over the entire population is called the fixation probability [3]–[8]. Fixation probability quantifies the likelihood of propagation of a single mutant in the population. When different types of individuals have the same fitness value, the resulting evolutionary dynamics are called neutral evolutionary dynamics. In this case, it is well known that the fixation probability of a single mutant on the complete graph is the reciprocal of the population size [8].

In reality, individuals do not necessarily interact with everyone in the population. They have relationships with some individuals, but not with others. This fact leads to the idea of complex contact networks of individuals. Neutral evolutionary dynamics such as voter-type models in complex networks have been extensively studied (e.g. [9]–[12]). It has been shown that the fixation probability of a single mutant in undirected and unweighted networks depends on the degree of the initially invaded node as well as update rules [3]–[5]. Edges in many real networks, however, have directionality. Examples of real networks include social networks in which a directed edge is drawn from the actor to the recipient of grooming behavior of rhesus



Figure 1. Fixation of type A or type B after the introduction of a type A mutant.

monkeys [14]. Other examples include e-mail social networks [13, 15], and ecological networks, in which the heterogeneity of parameters such as habitat size [16] and geographical biases such as wind direction [17] and reverine streams [18] are exemplary sources of directionality. Moreover, the edges in real networks are generally weighted [19]. This concept has been nicely introduced in a seminal paper on evolutionary dynamics on graphs [3].

In this study, we investigate the dependence of the fixation probability of a single neutral mutant in general directed (or weighted) networks on its initial location and on update rules. We study three major update rules that were introduced in [4, 5]. Our results are remarkably different from those obtained in the case of undirected networks in which the fixation probability of a mutant is determined by the degree of the node where the mutant is initially placed. In the case of directed networks, the fixation probability crucially depends on the global structure of networks. The difference between directed and undirected networks is striking, especially in the case of modular, spatial *or* degree-correlated networks.

2. Model

Consider a population of N individuals comprising two types of individuals—type A and type B. Let the fitness of type A and type B individuals be r and 1, respectively. In this study, we mainly focus on the case r = 1, which corresponds to neutral competition between A and B. The structure of the population is described by a directed graph $G = \{V, E\}$, where $V = \{v_1, \ldots, v_N\}$ is a set of nodes, and E is a set of edges, i.e. v_i sends a directed edge to v_j if and only if $(v_i, v_j) \in E$. Each node is occupied by an individual of either type. The fitness of the individual at node v_i is denoted by $f_i \in \{r, 1\}$. Each directed edge $(v_i, v_j) \in E$ is endowed with its weight w_{ij} , which represents the likelihood with which the type of individual at v_i is transferred to v_j in an update step. We set $w_{ij} = 0$ when $(i, j) \notin E$.

Consider the introduction of a single mutant of type A at node v_i in a population of N - 1 residents of type B. Then, type A either eventually fixates, i.e. takes over the entire population, or becomes extinct (i.e. fixation of B), as schematically shown in figure 1. We are concerned with the fixation probability of type A, which is denoted by F_i . When needed, we refer to any one of the three update rules introduced below by the superscript of F_i , such as F_i^{LD} and F_i^{IP} . Throughout the paper, we assume that G is strongly connected. A network is strongly connected if there is at least one directed path between any ordered pair of nodes. If G is not strongly connected, we can find two nodes v_i and v_j such that there is no direct path from v_i to v_j . In this case, the fixation probability F_i is always zero because the individual at v_j is never replaced by the mutant initially located at v_i [3]. Therefore, it is sufficient to investigate fixation probabilities in the most upstream strongly connected.



Figure 2. Schematics of three update rules.

3. Results

In this section, we analytically obtain a system of linear equations that gives the fixation probabilities of mutants at individual nodes for the three update rules (link dynamics (LD), invasion process (IP) and voter model (VM)) [4, 5]. An update event in the three rules is schematically shown in figure 2. Then, we compare the analytical results with numerical results obtained for various directed networks.

3.1. Link dynamics (LD)

Firstly, we consider the LD [4, 5]. In this case, one directed edge is selected for reproduction in each time step; the edge $(v_i, v_j) \in E$ is chosen with probability $f_i w_{ij} / \sum_{k,l} f_k w_{kl}$ for the type of individual at v_i to replace the type of individual at v_j . Individuals with larger fitness values and larger outgoing edge weights are more likely to reproduce than those with smaller fitness values and outgoing edge weights. Thus, the selection process acts on birth events. Alternatively, the selection process can be assumed to act on death events, and the edge (v_i, v_j) is chosen with probability $(w_{ij}/f_j)/(\sum_{k,l} w_{kl}/f_l)$ for the type at v_i to replace that at v_j . This implies that individuals with smaller fitness values and larger incoming edge weights are more likely to die than those with larger fitness values and smaller incoming edge weights. In fact, the fixation probability F_i^{LD} ($1 \le i \le N$) is the same under these two interpretations.

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We consider the case r = 1 (hence, $f_i = 1, 1 \le i \le N$) analytically. Suppose that a single mutant of type A invades v_i . The fixation probability is given by F_i^{LD} . By considering the next update event, we can recalculate F_i^{LD} as follows. With probability $w_{ij} / \sum_{k,l} w_{kl}$, the edge $(v_i, v_j) \in E$ is selected for reproduction. Then, type A individuals occupy v_i and v_j . We denote by $F_{\{i,j\}}^{\text{LD}}$ the fixation probability $w_{ji} / \sum_{k,l} w_{kl}$, the edge $(v_j, v_j) \in E$ is selected. Then, type A individuals are initially located at v_i and v_j but nowhere else. With probability $w_{ji} / \sum_{k,l} w_{kl}$, the edge $(v_j, v_i) \in E$ is selected. Then, type A becomes extinct, and type A will not fixate. With the remaining probability $\sum_{k \neq i, l \neq i} w_{kl} / \sum_{k,l} w_{kl}$, the configuration of type A and type B individuals does not change. Therefore, we obtain

$$F_{i}^{\text{LD}} = \sum_{j} \frac{w_{ij}}{\sum_{k,l} w_{kl}} F_{\{i,j\}}^{\text{LD}} + \frac{\sum_{j} w_{ji}}{\sum_{k,l} w_{kl}} \times 0 + \frac{\sum_{k \neq i, l \neq i} w_{kl}}{\sum_{k,l} w_{kl}} F_{i}^{\text{LD}}.$$
 (1)

To prove $F_{\{i,j\}}^{\text{LD}} = F_i^{\text{LD}} + F_j^{\text{LD}}$, consider for now N neutral types labeled 1, 2, ..., N that are initially placed at v_1, v_2, \ldots, v_N , respectively. On a finite graph G, one of the N types fixates eventually. The probability that type i or j fixates is given in two ways: $F_{\{i,j\}}^{\text{LD}}$ and $F_i^{\text{LD}} + F_j^{\text{LD}}$. This ends the proof. Using $F_{\{i,j\}}^{\text{LD}} = F_i^{\text{LD}} + F_j^{\text{LD}}$, we rewrite equation (1) as

$$\sum_{j} w_{ij} F_j^{\text{LD}} = F_i^{\text{LD}} \sum_{j} w_{ji}.$$
(2)

This is a system of linear equations giving F_i^{LD} . We note that F_i^{LD} can be interpreted as the reproductive value of the individual at node v_i [20, 21].

Equation (2) can be derived more rigorously via the dual process [1, 2, 7, 22]. Intuitively speaking, the dual process of a stochastic process is another stochastic process in which the time of the original process is reversed. The direction of edges in the dual process is opposite to that in the original process. By considering the dual process, we can understand the tree of family lines in the original process, which is called genealogy. When we go backward in time, two individuals sometimes 'collide' in the dual process. Such an event is called coalescence. In terms of the original process, coalescence corresponds to two individuals sharing a common ancestor. After two individuals coalesce in the dual process, they behave as a single individual, representing a single family line. As far as the fixation probability is concerned, LD with r = 1 is equivalent to the continuous-time stochastic process in which each edge $(v_i, v_j) \in E$ is selected for reproduction at the Poisson rate w_{ii} . Then, the dual process of LD is the continuous-time coalescing random walk on the network with all edges reversed, with a random walker initially located on every node [1, 2]. A coalescing random walk is defined as follows. Consider a walker at v_i moving to v_i at the Poisson rate w_{ii} . If there is another walker at v_i , the two walkers coalesce into one at v_i and thereafter behave as a single random walker. On a finite graph G, the N walkers eventually coalesce into one, which is consistent with the fact that the ancestors of all individuals are the same in the end. Then, F_i^{LD} is the stationary density of the single random walker at v_i , which is given by equation (2). As G is strongly connected, the random walk on G with all edges reversed defines an irreducible Markov chain. Because F_i^{LD} is the stationary density of this irreducible Markov chain, equation (2) with constraints $\sum_{i=1}^{N} F_i^{\text{LD}} = 1$ and $F_i^{\text{LD}} \ge 0$ always has a unique strictly positive solution.

The calculation of F_i^{LD} from equation (2) by using a standard method such as the Gaussian elimination requires $O(N^3)$ computation time. However, because relevant large networks are usually sparse, carrying out the Jacobi iteration may take much less time. The convergence of this iteration to \mathbf{F}^{LD} is guaranteed by the Perron–Frobenius theorem [23].

In undirected graphs, $w_{ij} = w_{ji}$ holds. Therefore, $F_i^{\text{LD}} = 1/N$ solves equation (2), giving a result previously reported in [4, 5]. In the case of weighted or directed networks, the complexity of equation (2) implies that F_i^{LD} is not always determined by the local characteristics of node v_i but is affected by the global structure of the networks.

Next, we argue that the mean-field (MF) approximation is not useful in most cases. Consider unweighted, but possibly directed, networks such that $w_{ij} = 1$ if $(v_i, v_j) \in E$ and $w_{ij} = 0$ otherwise. Let $k_i^{\text{in}}(k_i^{\text{out}})$ be the indegree (outdegree) of v_i , and we set

$$\bar{F}^{\rm LD} = \frac{1}{N} \sum_{i=1}^{N} F_i^{\rm LD}.$$
(3)

The relation $\sum_{j} w_{ji} = k_i^{\text{in}}$ combined with the MF approximation $\sum_{i} w_{ij} F_j^{\text{LD}} \approx \sum_{i} w_{ij} \bar{F}^{\text{LD}} = k_i^{\text{out}} \bar{F}^{\text{LD}}$, yields

$$F_i^{\rm LD} \propto \frac{k_i^{\rm out}}{k_i^{\rm in}}.\tag{4}$$

Equation (4) indicates that a large k_i^{out} aids in the dissemination of the type at v_i and a small k_i^{in} inhibits the replacement of the type at v_i by the type at other nodes.

However, the MF approximation deviates from the correct F_i^{LD} in many cases. As an example, consider the largest strongly connected component of a directed and unweighted e-mail social network [13] with N = 9079 nodes and $\langle k \rangle = \langle k^{\text{in}} \rangle = \langle k^{\text{out}} \rangle = 2.62$, where $\langle \cdot \rangle$ denotes the average over the nodes. In this case, F_i^{LD} (indicated by the circles in figure 3(c)) does not agree with the normalized $k_i^{\text{out}}/k_i^{\text{in}}$ (indicated by the line). This is mainly because the indegree and outdegree of the same node in this network are positively correlated and because this network presumably has a nontrivial global structure. Actually, the Pearson correlation coefficient (PCC) for the N pairs $(k_i^{\text{in}}, k_i^{\text{out}}), 1 \leq i \leq N$, defined by

$$\frac{\frac{1}{N}\sum_{i=1}^{N} \left(k_{i}^{\text{in}}k_{i}^{\text{out}} - \langle k \rangle^{2}\right)}{\sqrt{\frac{1}{N}\sum_{i=1}^{N} \left(k_{i}^{\text{in}} - \langle k \rangle\right)^{2}} \sqrt{\frac{1}{N}\sum_{i=1}^{N} \left(k_{i}^{\text{out}} - \langle k \rangle\right)^{2}}}$$
(5)

is equal to 0.40. Networks in which degrees of adjacent nodes are correlated also show considerable discrepancies between the MF approximation and the numerical results. On the other hand, in the case of undirected networks, our result $F_i^{\text{LD}} = 1/N$ holds true in the presence of degree correlation of any kind, which is consistent with previously obtained numerical results [5].

Next, we examine the fixation probability in an asymmetric small-world network constructed from a ring of N = 5000 nodes. This network is a directed version of the Watts–Strogatz small-world network [24]. Each node of this network tentatively sends directed edges to five nearest nodes along both sides. Then, 2500 out of 10N = 50000 directed edges are rewired so that their two ends are randomly and independently selected from the N nodes, excluding self-loops and pre-existing edges. The correlation between the in- and out-degrees of the same node is negligible, with the PCC for the pairs (k_i^{in}, k_i^{out}) , defined by equation (5), being equal to -0.021. The degrees of adjacent nodes v_i and v_j conditioned by the existence of the edge (v_i, v_j) [25, 26] are also uncorrelated by the definition of the model. Nevertheless, the MF approximation is not effective in predicting the actual F_i^{LD} , as shown in figure 3(d). This discrepancy persists even for large N, because the directed small-world network does not render



Figure 3. F_i^{LD} (i.e. the fixation probability under LD) for a single mutant initially at node v_i in (a) an asymmetric random network with N = 5000, (b) an asymmetric scale-free network with N = 5000, (c) the largest strongly connected component of an e-mail social network with N = 9079 and (d) an asymmetric small-world network with N = 5000. The normalized $k_i^{\text{out}}/k_i^{\text{in}}$ is equal to $(k_i^{\text{out}}/k_i^{\text{in}})/\sum_{j=1}^N (k_j^{\text{out}}/k_j^{\text{in}})$. The lines represent the MF ansatz $F_i^{\text{LD}} = (k_i^{\text{out}}/k_i^{\text{in}})/\sum_{j=1}^N (k_j^{\text{out}}/k_j^{\text{in}})$.

 F_i^{LD} of adjacent nodes independent of each other. In contrast, F_i^{LD} in undirected small-world networks is completely determined by the MF ansatz indicated by the line in figure 3(d).

In contrast to these networks, figure 3(a) shows that the MF relation $F_i^{\text{LD}} \propto k_i^{\text{out}}/k_i^{\text{in}}$ roughly holds for a directed random graph with N = 5000. We generate a directed random graph by connecting each ordered pair of nodes (v_j, v_j) with probability $2\langle k \rangle/(N-1)$, so that $\langle k^{\text{in}} \rangle = \langle k^{\text{out}} \rangle = \langle k \rangle$. In this network, degree correlation and macroscopic network structure are both absent, which enables the application of the MF approximation. Even in this network, however, the MF approximation is not exact because, as equation (2) predicts, F_i^{LD} of nearby nodes are positively correlated. Following [26], we measure the correlation, or the assortativity, of F_i^{LD} by the PCC for the pairs $(F_i^{\text{LD}}, F_j^{\text{LD}})$ for $(v_i, v_j) \in E$ defined by

$$\frac{\frac{1}{N}\sum_{(i,j)\in E} \left(F_i^{\text{LD}}F_j^{\text{LD}} - \left(\bar{F}^{\text{LD}}\right)^2\right)}{\frac{1}{N}\sum_{(i,j)\in E}^N \left(F_i^{\text{LD}} - \bar{F}^{\text{LD}}\right)^2},\tag{6}$$

where \bar{F}^{LD} is defined by equation (3). The value of the PCC turns out to be slightly but significantly positive (mean \pm standard deviation based on 100 network realizations is equal to 0.0834 ± 0.0057). The discrepancy is also significant for a large N, unless the mean degree $\langle k \rangle$ is large.

The results obtained for random networks extend to the case of scale-free networks without degree correlation. We generate a directed scale-free network by setting the degree distribution to be $p(k) \propto k^{-3}$ for $k \ge \langle k \rangle/2$ and p(k) = 0 for $k < \langle k \rangle/2$, thereby generating k_i^{in} and k_i^{out} ($1 \le i \le N$) independently according to p(k), and randomly connecting the nodes using the Molloy–Reed algorithm [27]. Figure 3(b) indicates that the MF approximation roughly explains the numerically obtained fixation probability.

It is noted that the PCC for the pairs $(F_i^{\text{LD}}, F_j^{\text{LD}})$ for $(v_i, v_j) \in E$ is small for the asymmetric scale-free network (= 0.0395 ± 0.0056), large for the asymmetric small-world network (= 0.7888 ± 0.0165) and small for the e-mail social network (= 0.0420).

3.2. Invasion process (IP)

Next, consider the IP [4, 5]. In the IP, selection acts on birth. In each time step, v_i is first selected for reproduction with probability $f_i / \sum_k f_k$, where $f_i \in \{r, 1\}$ is the fitness of the type at node v_i . Then, with probability $w_{ij} / \sum_l w_{il}$, the type at v_i replaces that at v_j . Consequently, the probability that the edge $(v_i, v_j) \in E$ is used for reproduction in an update step is equal to $f_i w_{ij} / (\sum_k f_k \sum_l w_{il})$. On the complete graph, IP is the same as the standard Moran process [6]. For an arbitrary *r*, the IP is mapped to the LD with the rescaled edge weight $w'_{ij} = w_{ij} / \sum_l w_{il}$. Therefore, from equation (2), the fixation probability for r = 1 is the solution to

$$\sum_{j} \frac{w_{ij}}{\sum_{l} w_{il}} F_j^{\text{IP}} = F_i^{\text{IP}} \sum_{j} \frac{w_{ji}}{\sum_{l} w_{jl}}.$$
(7)

In the case of undirected unweighted networks, $F_i^{\text{IP}} \propto 1/k_i$ solves equation (7), giving a previously obtained result [4, 5, 22]. In the case of directed unweighted networks, applying the MF approximation to equation (7) yields

$$F_i^{\rm IP} = \frac{\sum_{j,(i,j)\in E} F_j^{\rm IP}/k_i^{\rm out}}{\sum_{j,(j,i)\in E} 1/k_j^{\rm out}} \approx \frac{(\rm const)}{k_i^{in}}.$$
(8)

The numerical results for the asymmetric random graph that is used for obtaining the results shown in figure 3(a) are shown in figure 4(a). The relation $F_i^{\text{IP}} \propto 1/k_i^{\text{in}}$ (the line in figure 4(a)) is roughly satisfied. Similar to the case of LD, some deviation persists in the case of random networks even with a large N. In contrast, figure 4(b) indicates that the actual F_i^{IP} in the scalefree network deviates considerably from equation (8), mainly because of the discreteness of k_i^{in} for a small integer k_i^{in} . Similar to the case of LD, the discrepancy between equation (8) and the exact F_i^{IP} is also large in the case of the e-mail social network (figure 4(c)) and the asymmetric small-world network (figure 4(d)). In addition to the degree correlation or global structure of networks, the discreteness of $1/k_i^{\text{in}}$ for a small k_i^{in} causes further deviation, as shown in figures 4(c) and (d).

3.3. Voter model (VM)

We examine a third update rule, the so-called VM [4, 5]. In the VM, we first eliminate the type at one node v_j with probability $f_j^{-1} / \sum_k f_k^{-1}$. Then, with probability $w_{ij} / \sum_l w_{lj}$, the type at v_i replaces that at v_j . The probability that the edge $(v_i, v_j) \in E$ is used for reproduction in an update step is equal to $f_j^{-1} w_{ij} / (\sum_k f_k^{-1} \sum_l w_{lj})$. For general r, the VM is mapped to the



Figure 4. F_i^{IP} in (a) an asymmetric random network, (b) an asymmetric scalefree network, (c) an e-mail social network and (d) an asymmetric small-world network. The normalized $1/k_i^{\text{in}}$ is equal to $(1/k_i^{\text{in}})/\sum_{j=1}^N (1/k_j^{\text{in}})$. The lines represent the MF ansatz $F_i^{\text{IP}} = (1/k_i^{\text{in}})/\sum_{j=1}^N (1/k_j^{\text{in}})$.

LD with the rescaled edge weight $w'_{ij} = w_{ij} / \sum_l w_{lj}$. Consequently, from equation (2), F_i^{VM} for r = 1 is given by

$$\sum_{j} \frac{w_{ij}}{\sum_{l} w_{lj}} F_{j}^{\rm VM} = F_{i}^{\rm VM} \sum_{j} \frac{w_{ji}}{\sum_{l} w_{li}} \quad (=F_{i}^{\rm VM}).$$
(9)

In the case of undirected networks, $F_i^{VM} \propto k_i$ solves equation (9), recovering a previously obtained result [4, 5, 22]. The MF approximation yields

$$F_i^{\rm VM} = \frac{\sum_{j,(i,j)\in E} F_j^{\rm VM}}{k_i^{\rm in}} \approx (\text{const}) \times k_i^{\rm out}.$$
(10)

In the case of the random network and the uncorrelated scale-free network, the numerical results shown in figures 5(a) and (b), respectively, support the rough validity of equation (10). However, this naive ansatz deviates from the actual F_i^{VM} for the e-mail social network (figure 5(c)) and the asymmetric small-world network (figure 5(d)). This situation is similar to that observed in the case of LD.

The fixation probability F_i^{VM} on a graph G is equivalent to the PageRank of node v_i of the graph G', where G' is constructed by reversing all edges of G. The PageRank measures the number of directed edges a node, such as a webpage, receives from other important nodes as exclusively as possible [28]–[30]. If we neglect some minor technical treatments that are



Figure 5. F_i^{VM} in (a) an asymmetric random network, (b) an asymmetric scalefree network, (c) an e-mail social network and (d) an asymmetric small-world network. The normalized k_i^{out} is equal to $k_i^{\text{out}} / \sum_{j=1}^N k_j^{\text{out}}$. The lines represent the MF ansatz $F_i^{\text{VM}} = k_i^{\text{out}} / \sum_{j=1}^N k_j^{\text{out}}$.

necessary for the practical implementation, the PageRank F_i^{PR} of v_i is defined by

$$\sum_{j} \frac{w_{ji}}{\sum_{l} w_{jl}} F_{j}^{\text{PR}} = \lambda F_{i}^{\text{PR}},\tag{11}$$

where λ is the largest eigenvalue of the eigen equation (11). If many edges are directed to v_i , there are many positive terms (i.e. $w_{ji} > 0$) on the lhs of equation (11), and they contribute to the PageRank of v_i on the rhs. If v_i receives an edge from v_j whose outdegree is small or whose PageRank is large, $w_{ji}/\sum_l w_{jl}$ or F_j^{PR} is large. Each of these factors also increases F_i^{PR} . A strongly connected network yields $\lambda = 1$, so that F_i^{PR} is the stationary density of the discrete-time simple random walk on the original graph G [23], [28]–[30]. Equation (11) with $\lambda = 1$ and with w_{ij} replaced by w_{ji} is identical to equation (9). In PageRank, nodes that receive many edges tend to be important, whereas the opposite is true in the case of the VM. F_i^{PR} is locally approximated by using k_i^{in} [31], which corresponds to the MF relation $F_i^{VM} \propto k_i^{out}$. However, F_i^{PR} in real web graphs often deviates from the relation $F_i^{PR} \propto k_i^{in}$ [32]. This implies that F_i^{VM} in real networks can also deviate from the MF approximation, which is consistent with our main claim.

3.4. Constant selection $(r \neq 1)$

When $r \neq 1$, the fitness value of type A and that of type B are different, so one type has a unilateral advantage over the other type. We call this situation 'constant selection'. In this case,



Figure 6. (a) F_i^{LD} for r = 4 plotted against F_i^{LD} for r = 1. (b) F_i^{IP} for r = 4 plotted against F_i^{IP} for r = 1. (c) F_i^{VM} for r = 4 plotted against F_i^{VM} for r = 1. We have used an asymmetric random network with N = 200.

the dual process of the evolutionary dynamics is the coalescing and branching random walk [2], which is difficult to handle analytically. Therefore, we carry out Monte Carlo simulations for r = 4 on a fixed asymmetric random graph with N = 200 and $\langle k \rangle = 10$. We calculate F_i^{LD} as a fraction of runs from 2×10^6 runs in which the single mutant with fitness r initially located at v_i (i.e. $f_i = r$ and $f_j = 1$, $j \neq i$) eventually occupies all nodes of the network. In figure 6(a), the numerically obtained F_i^{LD} for r = 4 is plotted against the exact solution of F_i^{LD} for r = 1 (equation (2)). Roughly speaking, F_i^{LD} for r = 4 monotonically increases with the exactly obtained F_i^{LD} for r = 1. We obtain similar results in the case of the IP (figure 6(b)) and VM (figure 6(c)). Therefore, the node from which a mutant is more likely to propagate throughout the population under the neutral dynamics (r = 1) also serves as a better invading node for mutants under constant selection ($r \in 1$). Thus, our results derived in the case of neutral selection is useful in predicting the order of the magnitude of fixation probabilities under constant selection.

3.5. Modular networks

Real networks are often more complex than degree-uncorrelated random, scale-free or smallworld networks. In particular, many networks are modular, i.e. they consist of several densely connected subgraphs termed modules, and each subgraph is connected to each other by a relatively few edges [33]. This is also the case for directed [34, 35] and weighted [36] networks.



Figure 7. (a) An example of a modular network, with w_{11} , w_{12} , w_{21} and w_{22} indicating edge weights when this network is coarse grained as a two-node network. (b) F_i^{LD} , (c) F_i^{IP} and (d) F_i^{VM} in an asymmetric modular network with N = 5000. The solid lines represent the MF ansatz (see the captions of figures 3–5 for details). The dashed lines represent the ansatz derived from the combination of the local degree and the module membership of the node (see equation (12) for the expression in the case of LD).

To intuitively understand the importance of the global structure of networks such as community structure in evolutionary dynamics, we generate a modular network [35], as schematically shown in figure 7(a), and study the fixation probability under neutrality, i.e. r = 1. We generate two modules $M_1 = \{v_1, \ldots, v_{N/2}\}$ and $M_2 = \{v_{N/2+1}, \ldots, v_N\}$ as two directed random graphs with N/2 = 2500 nodes and the mean degree $\langle k \rangle_M = 10$. Then, we randomly connect M_1 and M_2 by directed edges so that a node in M_1 (M_2) has $w_{1\rightarrow 2}\langle k \rangle_M$ ($w_{2\rightarrow 1}\langle k \rangle_M$) outgoing edges to the nodes in M_2 (M_1) on average. By setting $w_{1\rightarrow 2} = 0.04$ and $w_{2\rightarrow 1} = 0.01$, we obtain a network with N = 5000 and $\langle k \rangle = 10.25$. Note that the degree correlation is absent in this network. For the realized network, F_i^{LD} for r = 1 is shown in figure 7(b). Rather than the MF ansatz $\propto k_i^{\text{out}}/k_i^{\text{in}}$ (solid line), the module membership is the main determinant of F_i^{LD} . The upper and lower sets of points in figure 7(b) correspond to the nodes in M_1 and M_2 , respectively. The magnitude of F_i^{LD} in the two sets differ approximately by a factor of $w_{1\rightarrow 2}/w_{2\rightarrow 1} = 4$. The results obtained in the case of the IP and VM are similar, as shown in figures 7(c) and (d), respectively. There, gross connectivity among modules, not local degrees, principally determines F_i . A modified ansatz that combines the MF approximation and the multiplicative factor determined by the module membership of the node

$$F_i^{\text{LD}} \propto \begin{cases} w_{1 \to 2} k_i^{\text{out}} / k_i^{\text{in}}, & v_i \in M_1 \\ w_{2 \to 1} k_i^{\text{out}} / k_i^{\text{in}}, & v_i \in M_2 \end{cases}$$
(12)

fits the data well (the dashed lines in figure 7(b)). Similar approximations in which $k_i^{\text{out}}/k_i^{\text{in}}$ in equation (12) is replaced with $1/k_i^{\text{in}}$ and k_i^{out} (dashed lines in figures 7(b) and (c), respectively) roughly agree with the observed F_i^{IP} and F_i^{VM} .

To explain this result analytically, we presume that all nodes in a module are equivalent and have an identical fixation probability, \hat{F}_1 or \hat{F}_2 . In this manner, a network with two modules is reduced to a network with two nodes and self-loops. Equation (2) with N = 2 yields

$$\hat{F}_{1}^{\text{LD}} = \frac{w_{12}}{w_{12} + w_{21}},\tag{13}$$

$$\hat{F}_{2}^{\text{LD}} = \frac{w_{21}}{w_{12} + w_{21}}.$$
(14)

Because $w_{11} = w_{22} = \langle k \rangle_M$, $w_{12} = w_{1 \rightarrow 2} \langle k \rangle_M$ and $w_{21} = w_{2 \rightarrow 1} \langle k \rangle_M$, we obtain

$$\frac{\hat{F}_{1}^{\text{LD}}}{\hat{F}_{2}^{\text{LD}}} = \frac{w_{1 \to 2}}{w_{2 \to 1}},\tag{15}$$

which agrees with the numerical results. On the other hand, $k_i^{\text{out}}/k_i^{\text{in}}$ is equal to $(1 + w_{1 \rightarrow 2})/(1 + w_{2 \rightarrow 1})$ and $(1 + w_{2 \rightarrow 1})/(1 + w_{1 \rightarrow 2})$ for M_1 and M_2 , respectively. Both of these values are close to unity when $w_{1 \rightarrow 2}$, $w_{2 \rightarrow 1} \ll 1$, i.e. when the network is modular. Therefore, the MF approximation gives $\hat{F}_1^{\text{LD}}/\hat{F}_2^{\text{LD}} \approx 1$, which is different from our simulated results.

Similar calculations in the case of the IP yield

$$\hat{F}_{i}^{\text{IP}} = \frac{C_{i}^{\text{IP}}}{C_{1}^{\text{IP}} + C_{2}^{\text{IP}}}, \quad (i = 1, 2),$$
(16)

where

$$C_1^{\rm IP} \equiv \frac{w_{12}}{w_{11} + w_{12}},\tag{17}$$

$$C_2^{\rm IP} \equiv \frac{w_{21}}{w_{21} + w_{22}}.\tag{18}$$

Therefore, we obtain

$$\frac{\hat{F}_1^{\text{IP}}}{\hat{F}_2^{\text{IP}}} \approx \frac{w_{1\to 2}}{w_{2\to 1}} \tag{19}$$

when $w_{1\to 2}$, $w_{2\to 1} \ll 1$. However, the naive MF ansatz (equation (8)) yields $1/k_i^{\text{in}} = 1/(\langle k \rangle (1 + w_{2\to 1})) \approx 1/\langle k \rangle$ $(1 \le i \le N/2)$ and $1/k_i^{\text{in}} = 1/(\langle k \rangle (1 + w_{1\to 2})) \approx 1/\langle k \rangle$ $((N/2) + 1 \le i \le N)$. Then, $\hat{F}_1^{\text{IP}}/\hat{F}_2^{\text{IP}}$ would be approximately unity, which does not well explain the simulation results shown in figure 7(c).

In the case of the VM, we obtain

$$\hat{F}_{i}^{\rm VM} = \frac{C_{i}^{\rm VM}}{C_{1}^{\rm VM} + C_{2}^{\rm VM}},\tag{20}$$

where

$$C_1^{\rm VM} \equiv \frac{w_{12}}{w_{12} + w_{22}},\tag{21}$$

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$$C_2^{\rm VM} \equiv \frac{w_{21}}{w_{11} + w_{21}}.$$
(22)

When $w_{1\rightarrow 2}, w_{2\rightarrow 1} \ll 1$, we obtain

$$\frac{\hat{F}_1^{\text{VM}}}{\hat{F}_2^{\text{VM}}} \approx \frac{w_{1\to 2}}{w_{2\to 1}}.$$
(23)

However, the naive MF ansatz (equation (10)) yields $k_i^{\text{out}} = \langle k \rangle (1 + w_{1 \to 2}) \approx \langle k \rangle$ $(1 \le i \le N/2)$ and $k_i^{\text{out}} = \langle k \rangle (1 + w_{1 \to 2}) \approx \langle k \rangle$ $((N/2) + 1 \le i \le N)$. Then, $\hat{F}_1^{\text{VM}} / \hat{F}_2^{\text{VM}}$ would be approximately unity, which again does not explain the simulation results shown in figure 7(d).

In sum, for each update rule the community structure of networks has a strong impact on the fixation probability.

4. Conclusions

In summary, we obtained the general formulae for the fixation probability in directed and weighted networks. For each of the three different update rules, fixation probability is a solution to a system of linear equations. Fixation probability in undirected networks is completely determined by the local connectivity [4, 5] under neutrality. In contrast, in the case of directed degree-correlated, small-world, or modular networks, fixation probability is not only determined by the degree of the node that a mutant initially invades, and it deviates from the MF approximation to a large extent. Our results indicate that the global connectivity of networks has a significant effect on the fixation probability.

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